



Review article

Technology-Based Interventions to Reduce Sexually Transmitted Infections and Unintended Pregnancy Among Youth



Laura Widman, Ph.D. ^{a,*}, Jacqueline Nesi, M.A. ^b, Kristyn Kamke, M.S. ^a,
 Sophia Choukas-Bradley, Ph.D. ^c, and J. L. Stewart ^a

^a Department of Psychology, North Carolina State University, Raleigh, North Carolina

^b Department of Psychology and Neuroscience, University of North Carolina at Chapel Hill, Chapel Hill, North Carolina

^c Department of Psychology, University of Pittsburgh, Pittsburgh, Pennsylvania

Article history: Received November 10, 2017; Accepted February 12, 2018

Keywords: Adolescent sexual health; Sex education; Technology-based interventions; eHealth; mHealth; Digital health; Sexually transmitted diseases; HIV prevention

A B S T R A C T

Purpose: Technology-based interventions to promote sexual health have proliferated in recent years, yet their efficacy among youth has not been meta-analyzed. This study synthesizes the literature on technology-based sexual health interventions among youth.

Methods: Studies were included if they (1) sampled youth ages 13–24; (2) utilized technology-based platforms; (3) measured condom use or abstinence as outcomes; (4) evaluated program effects with experimental or quasi-experimental designs; and (5) were published in English.

Results: Sixteen studies with 11,525 youth were synthesized. There was a significant weighted mean effect of technology-based interventions on condom use ($d = .23$, 95% confidence interval [CI] [.12, .34], $p < .001$) and abstinence ($d = .21$, 95% CI [.02, .40], $p = .027$). Effects did not differ by age, gender, country, intervention dose, interactivity, or program tailoring. However, effects were stronger when assessed with short-term (1–5 months) than with longer term (greater than 6 months) follow-ups. Compared with control programs, technology-based interventions were also more effective in increasing sexual health knowledge ($d = .40$, $p < .001$) and safer sex norms ($d = .15$, $p = .022$) and attitudes ($d = .12$, $p = .016$).

Conclusions: After 15 years of research on youth-focused technology-based interventions, this meta-analysis demonstrates their promise to improve safer sex behavior and cognitions. Future work should adapt interventions to extend their protective effects over time.

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IMPLICATIONS AND CONTRIBUTION

This meta-analysis demonstrates that technology-based interventions can effectively promote sexual health among youth. Compiling data from 16 studies, technology-based interventions were shown to improve condom use, abstinence, safer sex knowledge, and safer sex attitudes and norms. Effects were robust across many factors (e.g., age, gender, and dose) but were stronger with shorter term follow-up.

Conflicts of Interest: The authors have no financial relationships relevant to this article to disclose.

Disclaimer: The content, interpretations, and conclusions of this study are those of the authors and do not necessarily reflect the views of the National Institutes of Health or the National Science Foundation.

* Address correspondence to: Laura Widman, Ph.D., Department of Psychology, North Carolina State University, Campus Box 7650, Raleigh, NC 27695.

E-mail address: lmwidman@ncsu.edu (L. Widman).

The risk of unplanned pregnancy and sexually transmitted infections (STIs), including HIV, is high among youth. Youth ages 13–24 comprise nearly half of the 20 million new STI cases and more than 20% of new HIV diagnoses each year in the United States [1,2]. Adolescents and young adults account for 45% of all new HIV infections worldwide [3]. Rates of unintended pregnancy are also elevated among youth, with girls ages 15–19 having higher rates of unintended pregnancy than girls and women in any other age group [4,5]. Further, complications from childbirth

are the second leading cause of death among 15- to 19-year-old girls worldwide [5].

Efforts to provide youth with sexuality education to reduce HIV/STIs and unintended pregnancy have been underway for decades [6,7]. Only recently, however, has a movement emerged toward utilizing technology-based platforms for sexual health program delivery [8–10]. Technology-based interventions have been alternatively referred to as eHealth, mHealth, digital media, or new media interventions, and typically utilize computers, smartphones, text messaging, and/or other web-based platforms. As digital tools have become increasingly accessible and sophisticated in their functionalities, many researchers have heralded the promise of technology-based interventions to improve sexual health [8–10]. Compared with traditional interventions, new media approaches may allow for broad reach at relatively low costs, improved fidelity during intervention delivery, greater privacy and comfort for teens learning about sensitive topics, and increased capacity for individually tailoring prevention messages [11,12]. Additionally, such tools may provide fruitful means for engaging today's digitally native youth, among whom the ubiquitous use of technology plays a central role in key developmental tasks [13,14].

The promise of new media tools has prompted a recent proliferation of technology-based interventions targeting sexual health among youth. Although recent systematic reviews have highlighted potential impacts and limitations of these interventions [15–17], little is known regarding their combined efficacy. Determining whether and through what mechanisms technology-based interventions promote youth sexual health is critical, as such knowledge can inform future interventions and justify the allocation of resources to their development.

To date, there have been only three published meta-analyses across adolescent and adult populations that provide insights into the efficacy of technology-based interventions for sexual health. The first meta-analysis focused on interventions to reduce HIV infection and included 12 randomized controlled trials published between 2002 and 2008 with participants of any age [18]. This study found a small but significant overall increase in condom use among participants who completed a technology-based intervention compared with those who were randomized to a control group, with an effect size that was comparable with similar in-person interventions (d across studies = .26). These effects were stronger when interventions included individualized tailoring (i.e., materials matched to the needs of specific participants) and had a greater number of sessions. The second meta-analysis focused exclusively on programs employing computer-based administration, finding that such interventions were effective in improving sexual health knowledge, sexual self-efficacy, safer sex intentions, and safer sexual behavior among adolescents and adults [19]. The third meta-analysis identified significant effects of new media interventions for increasing condom use and STI testing within nonclinical populations [20]. Importantly, this meta-analysis found that interventions produced the largest effect sizes when they targeted female adolescents [20].

Despite these preliminary findings, at least three critical gaps inhibit our ability to draw conclusions about the overall efficacy of technology-based sexual health programs for youth. First, there are currently no meta-analyses focused exclusively on youth, even though youth are at heightened risk of unplanned pregnancies and STIs [3,4,21,22] and are also some of the most frequent users of new media technologies [13,23]. Second, no meta-analyses have included abstinence as a behavioral outcome.

Abstinence can be a developmentally appropriate objective of comprehensive sexuality education—particularly in programs targeting early adolescent populations. Third, prior meta-analyses and systematic reviews of technology-based sexual health programs have only included studies published through 2014 [15,16,18–20,24]. Given the rapidly changing landscape of technology and the many new interventions introduced each year, an updated review of the literature is warranted.

Thus, the goal of this meta-analysis is to synthesize the growing literature on technology-based sexual health interventions among youth ages 13–24 and to determine their overall efficacy on two key behavioral outcomes: condom use and abstinence. Additionally, we will examine a number of secondary outcomes identified as important components of safer sexual decision-making within health behavior theories [25] that have guided many intervention efforts in this area. These outcomes include safer sex attitudes, norms, self-efficacy, behavioral intentions, and sexual health knowledge. Finally, we will examine whether characteristics related to the intervention recipients (age, gender, country) or the intervention design (delivery method, use of tailoring, program interactivity, follow-up duration) moderate the effectiveness of technology-based interventions on sexual health outcomes.

Methods

Search strategy

We conducted a comprehensive search of *Medline*, *PsycINFO*, and *Communication Source* databases to extract relevant studies published through May 2017. We used the following combination of key words, with asterisks used as “wild cards” to find any variations: (*adolescen** or *teen** or *youth* or *middle school* or *high school*) AND (*sexual health* or *safe* sex* or *sex* education* or *sexually transmitted disease* or *sexually transmitted infection* or *STD* or *STI* or *HIV* or *AIDS* or *pregnancy* or *reproductive health* or *condom** or *contracept** or *protected sex* or *unprotected sex*) AND (*intervention* or *program* or *education* or *trial*) AND (*technology* or *internet* or *web-base** or *computer-base** or *online* or *social media* or *social network** or *SNS* or *eHealth* or *mHealth* or *electronic health* or *mobile health* or *texting* or *text messag** or *digital media* or *new media* or *gaming* or *SMS* or *mobile phone* or *cell phone* or *phone app** or *Facebook* or *Twitter* or *Instagram* or *instant message* or *web 2.0* or *media 2.0*). Additional studies of potential relevance were located by examining prior reviews and meta-analyses [15–20,24]. We also examined the reference lists of all included articles to search for additional studies. This search produced an initial 1,932 scientific articles.

Selection criteria

Studies were included if they met the following criteria: (1) focused on youth between ages 13–24 (i.e., mean sample age 13–24 and no participant older than 29); (2) utilized technology as the primary mode of delivering an HIV/STI or pregnancy prevention intervention (studies that utilized technology but also included extensive in-person components were excluded [26,27]); (3) included a behavioral outcome measure of either condom use/unprotected sex (referred to as “condom use” in this paper; the effect sizes for unprotected sex were recoded so that the direction of effect always indicated greater protection) or abstinence/delayed intercourse (referred to as “abstinence”); (4) evaluated

program effects with an experimental or quasi-experimental design; (5) were published in English; and (6) provided sufficient statistics to calculate effect sizes. These selection criteria resulted in a final sample of 16 articles (see [Supplementary Figure S1](#)). In total, we calculated 16 independent effect sizes for condom use, 8 for abstinence, 6 for safer sex attitudes, 5 for norms, 5 for self-efficacy, 5 for behavioral intentions, and 9 for sexual health knowledge.

Data extraction

Two of the authors (K.K. and J.L.S.) independently coded the primary studies. The following data were abstracted: (1) demographic and sample characteristics; (2) intervention characteristics (e.g., type of technology used, number of sessions, and program interactivity); (3) methodological characteristics (e.g., length of follow-up and type of comparison group); and (4) type of outcome measurement(s). The mean percentage agreement across all coding categories was 88%. Discrepancies between coders were resolved through discussion with the first author (L.W.). Additionally, three of the authors (L.W., K.K., and J.L.S.) rated the 16 studies for quality and risk of bias using the Cochrane Collaboration's tool for assessing risk of bias (see [Supplementary Table S1](#)).

Calculation of effect sizes

The standardized mean difference, d , was used as the indicator of effect size (i.e., the treatment group and control group means divided by the pooled standard deviation). According to Cohen [28], effect size d can be interpreted as small (.20), medium (.50), or large (.80). When d 's were reported in an article, they were directly extracted. If d 's were not reported, other statistics that could be converted to d 's (e.g., summary statistics and odds ratios) were calculated using Comprehensive Meta-Analysis V2.0 [29]. When no statistics in the study could be converted to a d , study authors were contacted and appropriate data were requested. When more than one follow-up time point was reported, effect sizes were calculated based on data from the longest follow-up for which data were available. To ensure the consistency and interpretability of effect sizes, higher values always indicate that the technology-based intervention group performed better than the control group.

We used random effects meta-analytic procedures for the primary analyses across all independent effect sizes; this procedure allowed for the possibility of differing variances across studies [30]. The Q statistic and I^2 were used to examine whether significant heterogeneity existed among effect sizes. Effect sizes for hypothesized moderators were calculated along with their 95% confidence intervals, and those effect sizes were compared using the Q_b statistic. For these analyses, mixed effects models were utilized to allow for the possibility of differing variances across subgroups. These models employ random effects assumptions, while stratifying the effect sizes by fixed factors such as age and study dose [30]. Analyses were conducted using Comprehensive Meta-Analysis V2.0 [29].

Results

Study characteristics

[Table 1](#) provides a summary of the studies included in this meta-analysis, including sample characteristics and moderator

variables. A total of 11,525 participants (mean age = 18.42) were enrolled across 16 technology-based interventions for youth. The majority of studies were conducted in the United States ($k = 11$). Many studies used combined samples of boys and girls ($k = 9$); however, a few studies analyzed data from boys ($k = 3$) and girls ($k = 4$) independently.

The delivery of program content varied widely across different digital media technologies. Specifically, 5 programs were delivered exclusively via computer programs; 2 were delivered exclusively through Internet Web sites; 1 was delivered exclusively through texting; 1 was delivered exclusively through social media; and nearly half of programs (7 of 16) were delivered with more than one method, such as an Internet Web site with email follow-up [31,32,39] or a combination of text messaging and email delivery [38]. It was common for interventions to be interactive (i.e., accept input from the user as the program progressed; $k = 12$) and to utilize tailoring (i.e., include program components that were matched specifically to the user based on user characteristics; $k = 10$). The dose of intervention varied widely across studies, with $k = 5$ programs including 1–2 sessions and $k = 8$ programs including 7 or more sessions.

Primary outcomes

Condom use. Although individual effect sizes ranged from $d = -.07$ to .67, the weighted mean effect size for condom use across these studies was $d = .23$ (95% confidence interval [CI] [.12, .34]; $p < .001$), indicating that technology-based interventions have a small but significant protective effect on condom use behavior among youth (see [Figure 1](#)). To examine whether publication bias may have inflated the effect size of the interventions on condom use, a fail-safe N was calculated. This fail-safe N was 92, suggesting 92 nonsignificant studies would need to exist to reduce the effect size to a trivial level ($p > .05$).

Statistical testing indicated marginal heterogeneity among studies with regard to the condom use outcome ($Q = 24.27$, $p = .061$, $I^2 = 38.19$). Thus, we examined the potential impact of several moderating variables. As shown in [Table 2](#), there was only one marginally significant difference based on the duration of follow-up assessment, with shorter follow-ups (less than 6 months) producing stronger effects ($d = .32$, $p < .001$) than programs with follow-ups of longer duration (6 months or more; $d = .14$, $p = .093$). No significant differences were found by participant age, gender, study dose, program interactivity, program tailoring, or whether the study was conducted in the United States.

Abstinence. Individual study effect sizes for abstinence/delayed sex ranged from $d = .01$ to .65, with an overall weighted mean effect size for abstinence across studies of $d = .21$ (95% CI [.02–.40]; $p = .027$). This indicates that technology-based interventions have significant protective effects on abstinence over time (see [Figure 1](#)). To examine whether publication bias may have inflated the effect size of the interventions on abstinence, a fail-safe N was calculated. This fail-safe N was 18, suggesting 18 nonsignificant studies would need to exist to reduce the effect size to a trivial level ($p > .05$).

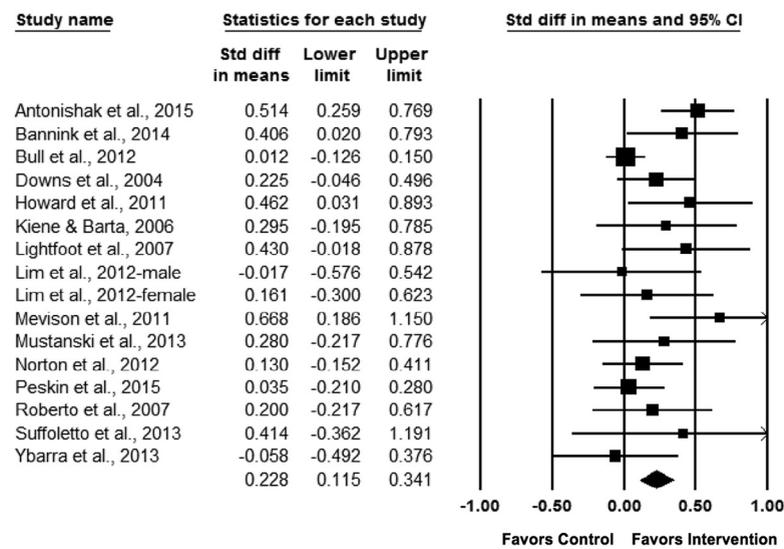
Statistical testing indicated significant heterogeneity among studies with regard to the abstinence outcome ($Q = 16.93$, $p = .018$, $I^2 = 58.65$); thus, we examined the potential impact of several moderating variables. As shown in [Table 3](#), the only significant moderator was the duration of follow-up. Similar to the effects on condom use, intervention effects on abstinence were

Table 1
Study characteristics

Authors	Sample and study description	Age, M (range)	Dose	Interactive	Tailoring	Follow-up (mo)	Outcome(s)
Antonishak et al. [31]	2,284 young women from the United States; race/ethnicity: 56% white, 17% black, 19% Hispanic; recruited online; intervention via Web site and email	22.3 (18–29)	High	Yes	No	12	Condom use
Bannink et al. [32]	1,702 adolescents from the Netherlands; 46% female; recruited in secondary schools; intervention via Web site and email	15.9 (15–16)	Low	No	Yes	4	Condom use
Bull et al. [33]	1,578 young adults from the United States; 65% female; race/ethnicity: 30% white, 35% black, 14% Hispanic; recruited through snowball sampling; intervention via social media	20.0 (16–25)	High	Yes	No	6	Condom use; self-efficacy; intentions
Downs et al. [34]	300 sexually active teenage girls from the United States; race/ethnicity: 15% white, 75% black, 10% mixed race; sexual orientation: 100% heterosexual; recruited in health center; intervention via computer	<i>nr</i> (14–18)	Moderate	Yes	No	6	Condom use; abstinence; knowledge
Howard et al. [35]	254 low-income teenage girls from the United States; race/ethnicity: 99% black; recruited in family planning clinics; intervention via Web site and computer	<i>nr</i> (“teen”)	Low	No	No	3–6	Condom use; knowledge
Kiene and Barta [36]	157 college students from the United States; 71% female; race/ethnicity: 81% white, 5% black, 6% Hispanic; sexual orientation: 98% heterosexual; recruited from psychology student participant pool; intervention via computer	18.9 (<i>nr</i>)	Low	Yes	Yes	1	Condom use; attitudes; norms; intentions; knowledge
Lightfoot et al. [37]	133 adolescents from the United States; 45% female; race/ethnicity: 49% black, 47% Hispanic; recruited from alternative high schools; intervention via computer	16 (14–18)	Moderate	Yes	Yes	3	Condom use; abstinence
Lim et al. [38]	994 young adults from Australia; 58% female; recruited from music festival; intervention via texting and email	19 (16–29)	High	No	No	12	Condom use; knowledge
Lou et al. [39]	1,337 adolescents from China; 45% female; recruited from high schools and universities; intervention via Web site and email	<i>nr</i> (15–19)	High	Yes	Yes	2.5	Abstinence; attitudes; norms; knowledge
Mevisen et al. [40]	171 young adults from the Netherlands; 61% female; sexual orientation: 100% heterosexual; intervention via Web site	20.8 (18–25)	Low	Yes	Yes	3	Condom use; attitudes; norms
Mustanski et al. [41]	102 sexually active young adult men from the United States; race/ethnicity: 25% white, 13% black, 46% Hispanic; sexual orientation: 100% men who have sex with men; recruited from HIV testing center; intervention via Web site and video game	21.3 (18–24)	Moderate	Yes	Yes	3	Condom use; attitudes; norms; self-efficacy; intentions; knowledge
Norton et al. [42]	198 sexually active college students from the United States; 70% women; race/ethnicity: 85% white; sexual orientation: 100% heterosexual; recruited from psychology student participant pool; intervention via computer DVD	18.6 (<i>nr</i>)	Low	No	No	2	Condom use
Peskin et al. [43]	1,571 adolescents from the United States; 59% female; race/ethnicity: 17% black, 74% Hispanic; recruited from urban middle schools in Texas; intervention via computer	14.3 (<i>nr</i>)	High	Yes	Yes	12	Condom use; abstinence; attitudes; self-efficacy; intentions; knowledge
Roberto et al. [44]	326 10th grade adolescents from the United States; 56% female; race/ethnicity: 97% white; recruited from rural high schools; intervention via computer	15.6 (<i>nr</i>)	High	Yes	Yes	2.5	Condom use; abstinence; attitudes; norms; self-efficacy; knowledge
Suffoletto et al. [45]	52 sexually active young adult women from the United States; race/ethnicity: 65% black, 6% Hispanic; sexual orientation: 100% heterosexual; recruited from emergency rooms; intervention via text messages	21.4 (18–25)	High	Yes	Yes	3	Condom use; abstinence
Ybarra et al. [46]	366 sexually active adolescents from Uganda; 16% female; recruited from high schools; intervention via Web site	16.1 (13–19)	Moderate	Yes	Yes	6	Condom use; abstinence

Sexual orientation is included in the sample description whenever it was reported in the text; many studies did not report the sexual orientation of their participants. Dose: high (seven or more sessions), moderate (three to six sessions), low (one to two sessions). *nr* = not reported.

CONDOM USE



ABSTINENCE

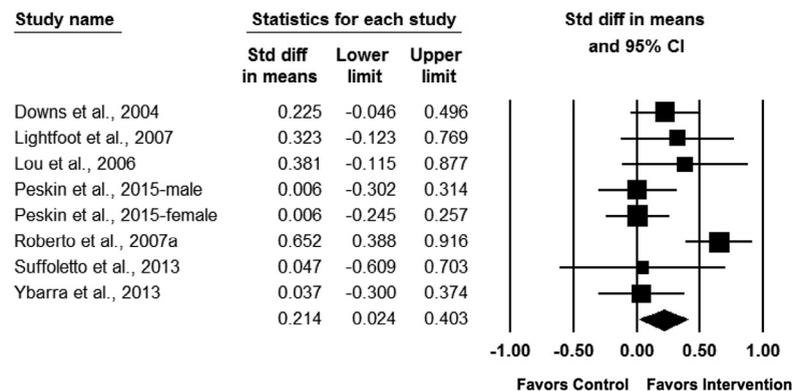


Figure 1. Forest plots for primary outcomes.

significantly stronger for short-term follow-ups (i.e., less than 6 months; $d = .45$, $p < .001$) than for long-term follow-ups ($d = .07$, $p = .321$). No significant differences were found by gender, age, intervention dose, tailoring, or country of study. Of note, all eight studies that included abstinence as a primary outcome were interactive interventions; thus, we could not examine the effects of interactive versus static programs for abstinence.

Secondary outcomes

We examined five secondary outcomes, including safer sex attitudes, social norms for safer sexual activity, self-efficacy, behavioral intentions to practice safer sex, and sexual health knowledge. As shown in Table 4, compared with control programs, technology-based interventions were efficacious in increasing sexual health knowledge ($d = .40$, $p < .001$), social norms for safer sexual activity ($d = .15$, $p = .022$), and safer sex attitudes ($d = .12$, $p = .016$) among youth. Compared with controls, there were no significant differences in the effect of technology-based interventions on safer sex intentions or perceived self-efficacy to engage in safer sexual behavior.

Data quality

The majority of study designs were rigorous (14 randomized controlled trials) and there was limited evidence of attrition bias, with retention rates across studies generally exceeding 80%. However, some uncertainty was noted in whether studies had utilized best practices for concealing the allocation of participants into study conditions. For eight studies, it was also unclear if selective reporting of outcomes could have been an issue since these studies had not been preregistered. However, overall data quality across studies was deemed to be high, and most studies were found to have a low risk of bias (see Supplementary Table S1).

Discussion

Results of the present meta-analysis, which synthesizes nearly 15 years of research on the development and evaluation of youth-focused technology-based interventions, highlight the great promise of these approaches to improve safer sex behaviors among youth. Pooling data from 16 studies with more than 11,000

Table 2
Weighted mean effect sizes for condom use by categorical moderator variables

Variable	<i>k</i>	<i>d</i>	95% CI	<i>p</i>	Between groups	
					<i>Q_b</i>	<i>p</i>
<i>Gender</i>					2.95	ns
Female only	4	.36	[.19, .53]	<.001		
Male only	3	.20	[−.14, .53]	.246		
Mixed gender	9	.17	[.03, .31]	.016		
<i>Age</i>					.33	ns
Sample mean age ≤18	9	.20	[.08, .31]	.001		
Sample mean age >18	7	.27	[.03, .51]	.025		
<i>Intervention dose</i>					.18	ns
Low (1–2 sessions)	5	.34	[.15, .52]	<.001		
Moderate (3–6 sessions)	4	.22	[.03, .41]	.025		
High (7+ sessions)	7	.17	[−.01, .35]	.070		
<i>Interactivity</i>					.00	ns
Interactive	11	.23	[.09, .38]	.002		
Static	5	.23	[.06, .41]	.002		
<i>Tailoring</i>					.07	ns
Used tailoring	9	.24	[.09, .39]	.002		
Did not use tailoring	7	.21	[.04, .39]	.018		
<i>Country</i>					.02	ns
U.S.	11	.22	[.09, .35]	.001		
Non-U.S.	5	.24	[−.02, .50]	.069		
<i>Follow-up</i>					2.73	.099
Less than 6 mo	9	.32	[.18, .47]	<.001		
6 mo or more	7	.14	[−.02, .31]	.093		

Mixed effects models are presented for moderator analyses.

CI = confidence interval; *d* = weighted mean effect size; *k* = number of studies; ns = not significant.

adolescents, this meta-analysis found a significant positive effect of technology-based sexual health interventions for improving two key sexual behaviors among youth: increasing consistent condom use (effect size *d* = .23) and delaying sexual activity (effect size *d* = .21). Across studies, technology-based programs also led to increases in sexual health knowledge, safer sex attitudes, and positive norms surrounding safer sexual activity. Importantly,

these effect sizes are comparable or even exceed the effects of in-person interventions, with generally small to moderate intervention effects noted for in-person programs (effect size *d* = .13 for condom use and *d* = .11 for abstinence in a meta-analysis by Johnson et al. [6]).

Technology-based interventions offer many possible benefits over traditional, face-to-face interventions [11,47,48]. Most

Table 3
Weighted mean effect sizes for abstinence by categorical moderator variables

Variable	<i>k</i>	<i>d</i>	95% CI	<i>p</i>	Between groups	
					<i>Q_b</i>	<i>p</i>
<i>Gender</i>					2.86	ns
Female only	2	.11	[−.11, .32]	.318		
Male only	2	.01	[−.27, .29]	.925		
Mixed gender	4	.36	[.05, .67]	.021		
<i>Age</i>					.26	ns
Sample mean age ≤18	7	.23	[.02, .43]	.030		
Sample mean age >18	1	.05	[−.61, .70]	.888		
<i>Intervention dose</i>					.06	ns
Low (1–2 sessions)	0	—	—	—		
Moderate (3–6 sessions)	3	.18	[−.01, .38]	.061		
High (7+ sessions)	5	.23	[−.08, .54]	.150		
<i>Interactivity</i>					—	—
Interactive	8	.21	[.02, .40]	.027		
Static	0	—	—	—		
<i>Tailoring</i>					.01	ns
Used tailoring	7	.21	[−.02, .44]	.069		
Did not use tailoring	1	.23	[−.05, .50]	.104		
<i>Country</i>					.10	ns
U.S.	6	.22	[−.01, .46]	.064		
Non-U.S.	2	.16	[−.16, .48]	.333		
<i>Follow-up</i>					7.00	.008
Less than 6 mo	4	.45	[.21, .69]	<.001		
More than 6 mo	4	.07	[−.07, .22]	.321		

Mixed effects models are presented for moderator analyses.

CI = confidence interval; *k* = number of studies; ns = not significant; *d* = weighted mean effect size.

Table 4
Effect sizes for safer sex-related supplemental outcomes

Variable	k	d	95% CI	p
Safer sex attitudes	6	.12	[−.02, .23]	.016
Safer sex norms	5	.15	[.02, .29]	.022
Safer sex self-efficacy	5	.07	[−.05, .19]	.249
Safer sex intentions	5	.06	[−.02, .14]	.158
Sexual health knowledge	9	.40	[.25, .55]	<.001

Random effects models are presented.

CI = confidence interval; d = weighted mean effect size; k = number of studies.

notably, these programs can be administered with high fidelity without extensive facilitator training. This may result in cost-effective programs that are capable of reaching a greater number of youth than are in-person programs. These approaches can also increase youths' openness to learning by providing a safe, controlled, and familiar environment to receive sexual health knowledge and skills. This may be especially salient for sexual minority youth who, in traditional, face-to-face interventions (e.g., school based), may receive sexual education that is stigmatizing, inaccurate, or irrelevant to their specific needs [49]. Further, technology-based programs offer ample opportunities for customizability, interactivity, and individual tailoring [11,47,48]. Technology-based interventions also have the potential to reach adolescents in resource-limited areas as worldwide access to mobile phones and Internet technologies is rapidly increasing among young people, even in low- and middle-income countries and rural areas [23,50].

The results suggest that the positive effects of technology-based interventions on sexual health may be robust across a number of different factors. Several potential moderating factors did not impact the efficacy of technology-based programs for youth, including the age and gender of participants and the intervention interactivity, tailoring, and dose. For example, although intervention dose varied from a single 15-minute session [35] to a full 12 months of content delivery [31,38], intervention dose did not influence effect sizes for condom use or abstinence across studies. This finding differs from a previous systematic review of in-person sexual health programs, which found that programs with higher duration and intensity produced the greatest treatment effects [51]. However, our findings are consistent with a more recent analysis of new media interventions [20], which found that the program duration did not influence the size of effects on condom use. Indeed, the effects of technology-based program dose on youth sexual health outcomes may be complex. On the one hand, the lack of association between intervention dose and outcomes could simply reflect the challenges of differentiating between the dose *intended* and the dose *received* within some technology-based programs [43]. On the other hand, this finding may reflect the growing potency of technology-based interventions. As technology has become increasingly available, interactive, and customizable in recent years, these interventions have clearly shown effectiveness, even at small doses. Future work could investigate whether such intervention impacts vary by how actively youth engage with technology on an everyday basis.

One crucial insight gleaned from this meta-analysis is that the effects of current technology-based programs tend to decrease over time. Specifically, for both condom use and abstinence, stronger effects were found in the short-term studies (i.e., 1- to 5-month follow-up) compared with studies that evaluated intervention

effects over 6 months or more. Adolescent sexual behavior is complex, and it is perhaps not surprising that program effects appear to diminish over time. The lack of long-term sustainability is a problem that has also plagued many programs delivered to youth in-person [52]. The knowledge and skills adolescents learn in specific programs may fail to adequately prepare them for long-term changes in their dating and sexual relationships and their evolving sexual interests and desires. Given the initial promise of technology-based programs, an exciting extension of this work would be to evaluate longer term, adaptive interventions utilizing Sequential Multiple Assignment Randomized Trial (SMART) designs [53]. These designs will allow investigators to examine the impact of added program components (e.g., booster sessions) over time to target the specific youth for whom treatment effects are diminishing. In fact, SMART designs may be particularly amenable to technology-based interventions, given unprecedented opportunities to automate assignment to intervention options (i.e., tailoring) based on participants' responses at critical decision points [54].

Limitations and future directions for intervention efforts

A number of important limitations in the studies included in this meta-analysis are worthy of future research attention. First, it is worth noting that there was substantial variation across studies in the way that the outcome measures were defined. For example, for both condom use and abstinence, some investigators focused exclusively on acts of vaginal sex [42,44,45] or anal sex [41], whereas others included acts of vaginal, anal, and oral sex in their definitions [43] or referred only to "sex" or "intercourse" without a definition [32,35]. Additionally, the timeline for measuring these behaviors differed. For example, some studies measured condom use at last sex [44,45], whereas others measured the frequency of condom use over a specified period of time [34,36]. Likewise, for the outcome measure of abstinence/delayed sex, some studies focused on whether youth had ever initiated sexual activity [43,44], whereas others focused on abstaining from sexual activity over a specified period of time [34]. Similar variation was observed in the definitions of the secondary outcomes under investigation. It is clear that the sexual health intervention literature lacks a "gold standard" for measuring sexual risk behavior and related attitudes and cognitions. This lack of consistency in measurement has been observed in other meta-analyses examining sexual risk behavior [6,18,55–57]. The measurement variation across studies could be obscuring our ability to detect the most precise estimates of intervention success, including which behaviors are most effectively addressed by technology-based interventions.

Second, only one of the included studies measured a biological outcome (i.e., Chlamydia infection [34]) and no programs assessed intervention effects beyond 1 year. These issues raise important questions about the long-term clinical impacts of these programs; however, they are not unique to technology-based programs. In fact, a recent systematic review of all school-based sexual health programs among youth in the United States [58] found only seven interventions that measured HIV/STI incidence or HIV/STI testing as outcomes. Further, among the 98 in-person interventions reviewed by Johnson et al. [6], the average outcome assessment occurred at 13 weeks. Future efforts to evaluate technology-based programs would be enhanced by data on HIV/STI testing and/or diagnosis among program recipients, as

well as long-term follow-up to measure the longitudinal impacts of these programs.

Another potentially fruitful avenue for efforts to improve adolescent sexual health outcomes will be to identify the specific, replicable behavioral change techniques that are utilized in technology-based interventions to enhance behavior change, as is increasingly being done for in-person interventions [59]. It will also be critical for researchers to identify common design elements that are *specific* to technology-based interventions going forward (for examples, see [48,60]). As technology rapidly evolves, recognizing common design features across programs (e.g., tailoring, online social support and communication, and automatic feedback) will be essential for comparing interventions delivered across a range of technological platforms. In this meta-analysis, we tested several possible technological design features that could serve as moderators of intervention effectiveness, such as whether the program was interactive and whether personal tailoring was utilized. However, there were too few studies to examine further design elements that may increase program impact, such as self-monitoring or personalized reminders [48]. These remain interesting avenues for future work as new technology-based programs continue to emerge [61,62].

As previously discussed, technology-based interventions may be a beneficial avenue for delivering nonstigmatizing, accurate, and relevant sexual health education to sexual minority youth. However, with the exception of one intervention targeting young men who have sex with men [41], no study included in this meta-analysis specifically reported tailoring intervention content to be inclusive of this population. Sexual minority youth experience disproportionate rates of sexual risk [63] and face a number of stressors that may impact their overall sexual health. Researchers should more explicitly consider the unique needs of sexual minority youth in the development and evaluation of sexual health programming, in addition to providing needed services at various levels of societal and individual intervention [64]. Technology-based approaches may be particularly valuable within this population, as sexual minority youth are more likely than heterosexual youth to use the Internet to look for sexual health information [65].

Future studies should also consider the potential moderating role of other important demographic factors, including race/ethnicity, gender, and socioeconomic status, given that these group comparisons were not possible in the present study. Sexual health disparities among racial and ethnic minorities have been well documented, with rates of reported STIs higher among some groups (e.g., black and Hispanic individuals) than white individuals [66]. Although a “digital divide” remains in terms of access and use of technology in the United States [67], recent statistics suggest that rates of smartphone access are comparable among Hispanic and white adolescents, and that black youth are *more* likely to have smartphones compared with other racial/ethnic groups [13]. Even among low-income youth (i.e., household income less than \$30,000), 91% of adolescents report that they access the Internet via a mobile device [13]. Thus, sexual health interventions delivered via technology may hold particular promise for reaching these at-risk groups.

Finally, there is a need for greater attention to the comparative effects of content delivered in-person versus through technology. Only one study in our meta-analysis compared a technology-based intervention with an in-person control [37]. However, the structure and quality of technology-based interventions may differ from in-person treatments in important ways,

including the possibility for anonymous communication (e.g., discussion boards), the provision of unique forms of social support and skills practice (e.g., using avatars or simulated person-to-person communication), and possibilities for different types of self-monitoring [48]. Additionally, with advancements in technology and incorporation of tailored, interactive features (e.g., “ask the expert”), technology-based interventions may offer the opportunity for personalized, rich communication, and support from intervention personnel that once was limited to in-person interventions [48,68]. However, interventions delivered in-person may increase the perceived credibility of information sources and provide a sense of accountability or connection [68]. Thus, although this meta-analysis suggests that technology-based interventions are effective compared with no-treatment or attention-matched controls, their *relative* efficacy in relation to traditional, in-person sexual health interventions remains a question for future work.

Summary and implications

After nearly 15 years of research on youth-focused technology-based interventions, this meta-analysis demonstrates the great promise of these approaches to improve the safer sex behavior and cognitions of youth. Pooling data from 16 studies with more than 11,000 adolescents and young adults, we found that technology-based interventions were more effective than control programs at improving condom use, abstinence, sexual health knowledge, safer sex attitudes, and safer sex norms. However, effects were stronger in studies with shorter term follow-ups than in studies with follow-ups of 6 months or more. Future work should focus on intervention adaptations and supplements that may extend their protective effects over time. Future work should also move beyond program development and into the realm of implementation science to ensure these programs are broadly disseminated for maximum impact [69].

Funding Sources

This research was supported in part by the following National Institutes of Health grants: R00 HD075654 and T32 MH018269. This work was also supported in part by a National Science Foundation Graduate Research Fellowship DGE-1144081.

Supplementary Data

Supplementary data related to this article can be found at <https://doi.org/10.1016/j.jadohealth.2018.02.007>.

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